AN INVESTIGATION OF THE ACCURACY OF THE PEARSON SELECTION FORMULAS

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and

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Office of Naval Research Contract 000-14-69C-0119 Melvin R. Novick, Principal Investigator



Educational Testing Service Princeton, New Jersey

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An Investigation of the Accuracy of the Pearson Selection Formulas

Melvin R. Novick and Dorothy T. Thayer

Introduction

The Pearson formulas for correcting correlation coefficients for restriction of range are based on crucial assumptions of linearity of regression and homoscedasticity of the error distributions. Some small studies, of which that of Rydberg (1963) is the most comprehensive, have previously been undertaken to determine the accuracy of these formulas. The general result found by Rydberg and others previously is that for small or moderate degree of selection the Pearson formulas are reasonably accurate but with some tendency to undercorrect. The present study was designed to investigate the accuracy of these formulas both for moderate and for extreme degrees of selection and to do so on many different types of variables. The unique feature of the present study is the very large sample sizes available for each of the data sets. With sample sizes of approximately 20,000 cases it is possible to perform extreme selection and still maintain relatively large sample sizes in the selected group. Thus investigations in these restricted subpopulations will not suffer from overly erratic fluctuation because of small sample sizes. While it is too much to say that in the case of extreme selection, we can treat the sample correlations within any extreme selection group as the true population correlation, still sampling variation should not appreciably distort our findings.

The central importance of a correction for restriction of range is apparent on noting that when comparing two tests, for example, a new test and an old test, as predictors of some criterion it is seldom possible to obtain criterion correlations for the total applicant group. Almost always selection must continue on

the basis of the old test and a valid comparison between tests cannot be made unless an accurate correction for restriction of range is available. If a correction is not made the general tendency will be to show the old test in a very unfavorable light, and thus to suggest the replacement of the old test with a new test when in fact such action is completely unwarranted.

In the present paper we restrict ourselves entirely to the case in which there exists a well defined explicit selection variable. Our purpose is to pin down as accurately as possible the range in which the Pearson formulas are acceptable both for explicit and incidental selection and to suggest other methods for cases in which they are not. A major problem in the application of range restriction corrections is the difficulty in isolating the actual selection variable. In most applications in which test scores are used for selection, they are not used on an exclusive basis so that, in fact, many other variables enter into selection. A popular way of "handling" this problem is to use a multivariate selection formula bringing in data on many incidental selection variables. The efficacy or even the logical justification for this approach has never been demonstrated. Moreover it should be clear that such a technique can be valid only to the extent that the simpler univariate and bivariate explicit and incidental selection formulas are valid when selection has, in fact, been explicit. Thus we are thrust back to the fundamental task of evaluating the simplest selection formulas. If we are to aspire to a personnel technology, as opposed to a personnel alchemy, we must be sure that popular corrections really do provide the needed corrections.

Description of Data Sets

Two major data sets were used in this study. The first of these was that used by Halpern to obtain norms for the Preliminary Scholastic Aptitude Test (PSAT) and the Academic Interest Measures (AIM). Halpern's data consisted of test scores on approximately 60,000 students in 180 schools. These students had taken the PSAT, the AIM and had completed a student questionnaire. The PSAT provided a verbal aptitude and a mathematical aptitude score. The AIM provided measures of interest in Biological Sciences, English, Fine Arts, Mathematics, Social Sciences, Secretarial, Physical Sciences, Foreign Languages, Music, Engineering, Home Economics, and Executive Occupations. Data from the student questionnaire were not used in the present study. The PSAT-AIM data consisted of scores on approximately 21,000 sophomores, 20,000 juniors and 18,000 seniors. It was decided that for the present study we would limit ourselves to the juniors. For this group data were available on 19,584 students. However data on many students were incomplete. Therefore for convenience it was decided to base our analysis only on those students who had complete scores on all PSAT and AIM scales. Data on 17,001 such students were available. The PSAT-V score is a scale score based on responses to the 70 PSAT-V items. Similarly PSAT-Q score gives a scaled score in the range 20-80 based on the marks on the 50 PSAT-Q items. Specifically, in each case the final score is obtained by taking the number of correct responses, subtracting a percentage of incorrect responses to get a formula score, and then linearly scaling into the interval 20-80. Each of the 12 AIM scale scores is based on responses to 16 keyed items of the AIM inventory. The reported scores are scaled so as to lie in the range 0-32.

Distributions of Test Scores

Table 1 gives the approximate means and standard deviations for the group on each of the two PSAT scales and 12 ATM scales. Table 2 gives the univariate

Insert Tables 1 and 2 about here

distributions of the PSAT-V and Q scores together with the percentage of the population at each score level and the cumulative percentage to that level.

A cursory inspection of Table 2 indicates that both PSAT distributions are positively skewed for this sample. This can be verified by noting that the coefficients of skewness of the distributions are .57 and .67 while a value of 0 would indicate a symmetric distribution. Since the mean of the PSAT-V scores is far below the center of possible values and nearly 7% of the scores are at the lowest attainable score, 20, it is clear that for the PSAT-V scale there is in fact some floor effect. The fact that a score of 80 was not attained on either scale indicates that no ceiling effect was present. On the whole then the two tests were somewhat difficult for the population of examiness. The coefficients of kurtosis were also computed and found to be -.26 and -.27 indicating that each of the distributions was platykurtic.

Table 3 gives the distributions of each of the AIM scales together with the computed coefficients of skewness and kurtosis. It is clear that for the most part we do not have either symmetric or mesokurtic distributions. We find

Insert Table 3 about here

for Home Economics, Secretarial, Foreign Language, and Executive scales a definite ceiling effect; thus, there is a tendency towards negative skewness and substantial platykurtosis.

A primary interest of this study was to determine the degree of linearity and homoscedasticity to be found typically among psychological variables. In order to conveniently investigate these aspects for the PSAT-Q and PSAT-V bivariate distribution it was decided to group the data into 21 small class intervals on each of the variables. The resulting bivariate plot is given in Table 4.

Insert Table 4 about here

Table 5 gives the means and standard deviations for each variable when the group is restricted to one of the class intervals on the second variable. In

Insert Table 5 about here

making the computations for this particular table, we have worked from the data in Table 4 and taken each person's score to lie at the midpoint of the class interval in which he falls. While the resulting computations for this table will have some degree of inaccuracy this should not be great because of the smallness of the class intervals. The obvious and important findings from this table are that the two regression lines tend to be reasonably linear except in the very extreme ranges but that the scedastic functions are not at all constant. Thus apparently one of the necessary assumptions of the use of the Pearson formulas is reasonably well satisfied except for extreme selection while the second assumption is not.

In order to facilitate processing of the PSAT data it was decided to further group the data into class intervals on each of the V and Q scales so that as nearly as possible each interval on each scale contains 10% of the population. Table 6 gives a bivariate plot of the PSAT-V and Q scores grouped into these class intervals and the cumulative percentages for each class interval.

Insert Table 6 about here

Table 7 gives the means and standard deviations for each variable when restricted subpopulations are defined by class intervals on the other variable.

Insert Table 7 about here

A cursory inspection of this table makes clear the relative acceptability of the linearity assumption and the complete unacceptability of the homoscedastic assumption. We do not give bivariate plots of the AIM scales with the PSAT scales since it is clear that there is a greater degree of nonnormality in the AIM scales. Thus the assumptions of linearity and homoscedasticity are even less likely to be satisfied.

The second data set used in this study was furnished by Dr. Milton H. Maier of the United States Army Behavioral Science Research Laboratory. Data were furnished on approximately 23,000 subjects; however, data on some subjects were incomplete. Only those subjects with complete information were used giving us 22,172 subjects. Due to technical difficulties 40 cases were lost during processing so that the majority of the results are reported for a total of 22,132 cases. Data on each subject consisted of 11 test scores on the Army Classification Battery, 14 test scores on the Army Differential MOS Battery, the MOS number of the training course to which the subject had been assigned and his final grade in that course.

The Army Classification Battery consists of 11 scales. The names of these scales and the number of items on which they are based are given in Table 8.

Insert Table 8 about here

Thirteen of the scales were taken from the Army Differential MOS Battery. The names of these scales and the numbers of items on which each is based are given in Table 9.

Insert Table 9 about here

Table 10 gives the means and standard deviations of each of the ACB and Differential MOS scales in the applicant group. Table 11 gives the bivariate

Insert Tables 10 and 11 about here

plot for ACB-V and ACB-A when the data have been grouped into 15 class intervals on each of the V and A scales. Table 12 gives the mean and standard deviation

Insert Table 12 about here

of the V and A scores when each of these variables has been restricted to one of the class intervals on the other variable. For these data the homoscedasticity assumption seems better satisfied than in the PSAT data. Coefficients of skewness and kurtosis were computed for both the V and A scales. Coefficients of skewness are -.14 and -.15 and the coefficients of kurtosis are -.36 and -.21.

Experimental Method

The first analysis used the PSAT-AIM data. We performed explicit selection on PSAT-V and assumed incidental selection on PSAT-Q and the 12 AIM variables. The strategy employed was to actually select on PSAT-V variable, determine the relevant correlations in the restricted population using the Pearson formula to correct for restriction of range and then to compare these adjusted values with

the actual correlations in the applicant group. Initial computations then involved determining means, standard deviations, and intercorrelations of the variables in the subpopulations defined by outselection on the left on PSAT-V in the groups 20-21, 20-25, 20-28, 20-34, 20-37, 20-41, 20-45, 20-51, and 20-56. In addition these computations were made for the applicant group, PSAT-V score 20-80.

The corrected correlation matrices and standard deviations were calculated assuming: there was explicit selection on variable X, PSAT-V; the variance of X, and the correlations and intercorrelations are known for the selected group and only the variance of X is available for the applicant group. The formulas used are given below. In these formulas small letters refer to the selected group, and capital letters refer to the applicant group.

$$R_{XY} = \frac{r_{xy}S_X}{\sqrt{s_x^2 - s_x^2 r_{xy}^2 + S_X^2 r_{xy}^2}}$$
(1)

$$R_{YZ} = \frac{s_{x}^{2}[r_{yz} - r_{xy}r_{xz}] + r_{xy}r_{xz}s_{x}^{2}}{\sqrt{[s_{x}^{2} - s_{x}^{2}r_{yy}^{2} + r_{xy}^{2}s_{x}^{2}][s_{x}^{2} - s_{x}^{2}r_{xz}^{2} + r_{xz}^{2}s_{x}^{2}]}}$$
(2)

$$s_{y} = s_{y}\sqrt{1 - r_{xy}^{2} + r_{xy}^{2} \frac{s_{x}^{2}}{s_{y}^{2}}}$$
 (3)

where R_{YZ} and r_{yz} are the correlations between two incidental selection variables,

 R_{XY} , r_{xy} and r_{xz} are the correlations of the explicit selection variable with an incidental selection variable,

 s_{χ} and s_{χ} are the standard deviations of the explicit selection variable, and s_{χ} and s_{χ} are the standard deviations of the incidental selection variables.

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A similar procedure was used with the Army data. Explicit selection was made on ACB-V called the X variable and incidental selection was assumed for the other seven ACB variables. Again the analysis assumed that the correlations and intercorrelations were known for the selected group and the variances of X were available for both the selected and applicant groups.

Comparison of the corrected correlations from restricted population with the values from the applicant group did not show the Pearson formulas in good light particularly when selection was at all severe. This is documented in the next section. In an attempt to discover a more generally useful correction, particular attention was given to the scedastic functions. This was done because it was found that the failure of this function to be constant was the primary violation of assumptions exhibited by both sets of data, though more so for PSAT data than the Army data. While several techniques were studied only two showed any promise and only these techniques are reported on here. These techniques involved discarding the assumption of constant error variance and using the assumption that the error variances have a general linear form. Attempts were then made to estimate the parameters of this linear relationship and thus to estimate the residual variance in the total population and to use this to obtain an improved correction for restriction of range.

Analysis of the PSAT-AIM Data

To evaluate the accuracy of the Pearson selection formulas with respect to the PSAT-AIM data, explicit selection was performed on PSAT-V with uccessive percentages in the selected group being approximately 10, 20, 30, 40, 50, 70, 80 and 90. The general pattern of results is illustrated in Table 13.

Insert Table 13 about here

The extrapolated correlations between PSAT-V and PSAT-Q are consistent with previous findings. The correlations (.36, .49, .58, .65, .73) in the successive groups with increasing percentages of selection are extrapolated to the values (.66, .68, .71, .73, .75). Since the true total group correlation is .75 this is certainly a clear and meaningful improvement, though, as found in previous studies, there is a tendency to undercorrect. Despite this substantial correction, however, one can question whether the correction is really adequate.

Suppose FSAT-V is the standard predictor of FSAT-Q and suppose that in current practice the selected group is 50% of the applicant group. Suppose further that a new predictor is being proposed and that this new predictor has a very low correlation with PSAT-V. (Actually this last assumption is most unlikely to occur in practice. We would be most fortunate if it did.) Then the restriction effect on the new predictor would be very small. In applications such as this an increase of .05 in the correlation coefficient would be considered a major advance, yet the Pearson extrapolated validity for the "old" test (PSAT-V) is .04 less than the actual total population value. Clearly in such a case there can be little justification for having any faith in the analysis. To compound the problem further one needs only note that for all other variables the typical result has been an overcorrection.

The fact that the correction works substantially less well with the AIM scales is a clear reflection of the sensitivity of the correction formula to the linearity and homoscedasticity assumptions. These scales are based on fewer items than are the PSAT scales, therefore asymptotic normality and hence linearity and homoscedasticity are much less evident.

It is also worth noting that there is a definite tendency for the correction formulas to be more accurate when the correlation in the applicant population

is substantial. When this correlation is near zero, the formulas seem to be of almost no value.

To evaluate the accuracy of the incidental selection formula we studied the correlations and extrapolated correlations between PSAT-Q and nine of the AIM variables. The results of this analysis are given in Table 14.

Insert Table 14 about here

Analysis of the Army Data

To evaluate the accuracy of the Pearson selection formulas with respect to the Army data explicit selection was performed on ACB-V with successive percentages in the selected group being approximately 10, 20, 40, 60, 80 and 90. The general pattern of results obtained can again be illustrated by looking at a few selected results given in Table 15.

Insert Table 15 about here

The extrapolated correlations between ACB-V and ACB-A, in this case, contradict previous findings. The correlations (.16, .27, .38, .47, .55, .58) in the successive groups with increasing percentages of selection are extrapolated to the values (.42, .60, .64, .65, .65, .64). Since the applicant group correlation is .60 this is again a clear and meaningful improvement. But in this case there has been a nontrivial overcorrection of even greater magnitude than the undercorrection in the PSAT-AIM data and again the correction with respect to all other variables has been an overcorrection.

To evaluate the accuracy of the incidental selection formula we examined

Table 16 which gives the correlations and extrapolated correlations between ACB-A

and six of the ACB scales. Here again the results were generally unsatisfactory.

Insert Table 16 about here

New Methods for Correcting Correlation Coefficients

Since the assumption of homoscedasticity of the error distribution does not appear to be satisfied for either data set, we attempted to find a procedure which would take into account the heteroscedasticity of errors. Using the linearity of regression assumption we have the following

$$R_{XY} = r_{XY} \frac{s_Y s_X}{s_X s_Y} . \qquad (4)$$

The standard deviations of the explicit and incidental selection variables, s_x and s_y , in the selected group and the standard deviation of the explicit selection variable, s_x , in the applicant group are known. The standard deviation of the incidental selection variable, s_y , in the applicant group has to be estimated to correct the correlation coefficient, r_{xy} , for restriction of range. Two methods of estimating s_y were attempted.

New method 1 used the analysis of variance breakdown of total variance into the sum of (1) average within-class variance and (1i) among-class variance.

$$\sigma^{2}(Y) = \varepsilon(\sigma^{2}(Y|X)) + \sigma^{2}(\varepsilon(Y|X)) . \qquad (5)$$

Thus, to estimate the variance of the incidental selection variable in the applicant group we had to estimate the expected value of the conditional

variances and the variance of the conditional means. Specifically this meant estimating $\sigma^2(Y|x)$ and $\varepsilon(Y|x)$ for those values of x in the rejected group.

The applicant group was divided into a selected group and a rejected group by selecting on an explicit selection variable. In the selected group, the conditional means and variances for the incidental selection variable were known for a number of intervals. These conditional means and variances were assumed to have a general linear form over these ordered intervals. By using least squares a straight line was fitted to the known conditional means of the incidental selection variable. The least squares estimates of the slope and intercept were used to obtain by extrapolation an estimate of the conditional mean for the incidental selection variable in the rejected group. The same procedure was used to estimate the conditional variance of the incidental selection variable in the rejected group. A weighted average of the known and estimated conditional variances of the incidental selection variable was used as a pooled estimate of the average conditional variance. The weights used were the number of persons in each class interval. Thus, an estimate of the first term in (5) was obtained.

The second term in (5) was estimated by using the relationship

$$\sigma^{2}\{\varepsilon(Y|x)\} = \varepsilon\{\varepsilon(Y|x)\}^{2} - \{\varepsilon(Y)\}^{2} . \tag{6}$$

An estimate for the average value of the incidental selection variable was calculated by pooling the known and estimated conditional means of the incidental selection variable. A similar procedure was used to estimate the average value of the squared conditional mean for the incidental selection variable. Hence, by using (6) an estimate of the variance of the conditional means was obtained.

The estimates of the average value of the conditional variance and the variance of the conditional means were combined using (5) to obtain an estimate for the total variance.

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This procedure for estimating the total variance of the incidental selection variable was used for each selected group except the first group. By using (4) the corrected correlation coefficients for each selected group were obtained. Tables 17 and 18 give these corrected correlation coefficients.

Insert Tables 17 and 18 about here

New method 2 assumed that the variances of the incidental selection variable in the selected groups had a general linear form. A straight line was fitted to the variances of the incidental selection variable in the selected groups by using least squares. The least squares estimates of the slope and intercept were used to estimate the variance of the incidental selection variable in the applicant group. This procedure for estimating the variance of the incidental selection variable in the applicant group was used for each selected group except the first since at least two points are needed to fit a straight line. The corrected correlation coefficients for each selected group were obtained by using (4); the values are given in Tables 17 and 18.

Summary

The results of this study strongly suggest that corrections for restriction of range are unsatisfactory even for moderate degrees of selection. Initial attempts to develop more sensitive techniques by relaxing the homoscedasticity assumption were not successful but further developments along these lines are

possible. A more promising approach would involve transforming variables, particularly the criterion variable, so as to achieve the required linearity and homoscedasticity.

Reference

Rydberg, S. (1963). Bias in Prediction. Stockholm: Almqvist and Wiksell.

Table 1
Means and Standard Deviations for PSAT and AIM Scales

Scale	Mean	SD
PSAT		
Verbal	36.c	11.2
Quantitative	38.1	11.3
AIM		
English	16.0	8.5
Music	15.0	8.9
Social Sciences	17.0	9.1
Mathematics	14.8	9.8
Physical Sciences	15.6	9.9
Engineering	16.5	9.8
Home Economics	19.2	9.3
Fine Arts	17.3	8.4
Biological Sciences	16.4	8.5
Secretarial	15.4	8.6
Foreign Languages	17.9	10.3
Executive	18.4	7.7

Table 2
Univariate Distributions for PSAT Scales

		PSAT-V			PSAT-Q	
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pct.
22122345667899123345567890123456784567845678	1158 1582 1442 1583 1573 1573 1573 1573 1573 1573 1573 157	62222.34322343333412222321111111100010. 818469883251823314575581199768548098707	6.97287526834445556664777778888889999999999999999999999999	134 2 134 2 134 2 134 3 134 3 137 3 138 3	0.000.8370.8315187563802363388320648651950 .88908370.8315187563802363388320648651950	0.1.2.4.7.9.4.9.1.8.5.0.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.4.7.4.2.0.3.4.9.1.8.5.0.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.4.7.4.2.0.3.4.9.1.8.5.0.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.6.4.7.4.2.0.3.4.9.1.8.5.0.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.6.4.7.4.2.0.3.4.4.9.1.8.5.0.6.9.7.7.9.2.8.1.3.1.9.1.3.3.9.3.1.7.0.1.0.6.6.4.7.4.2.0.3.4.4.4.5.4.4.4.5.4.4.4.5.4.4.4.5.4.4.4.5.4.4.4.5.4
59	112	0.7	97.0	148	0.9	95.4

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Table 2 (continued)

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	I	SAT-V			PSAT-Q	
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pet.
60	67	0.4	97.4	73	0.4	95.9
61	86	0.5	97.9	136	0.8	96.7
62	90	0.5	98.4	114	0.7	97.3
63	43	0.3	98.7	81	0.5	97.8
64	42	0.2	98.9	71	0.4	98.2
65	31	0.2	99.1	35	0.2	98.4
66	33	0.2	99.3	40	0.2	98.7
67	42	0.2	99.5	26	0.2	98.8
68	19 14	0.1	99.6	62	0.4	99.2
69	14	0.1	99.7	54	0.3	99.5
70	ננ	0.1	99.8	26	0.2	99.7
71	15	0.1	99•9	3	0.0	99.7
72	6	0.0	99.9	3	0.0	99.7
73)÷	0.0	99•9	30	0.2	99•9
74	1	0.0	99•9	10	0.1	99.9
75	2	0.0	100.0	5	0.0	100.0
76	7	0.0	100.0	7	0.0	100.0
77	Q	0.0	100.0	0	0.0	100.0
78	0	0.0	100.0	0	0.0	100.0
7 9	1	0.0	100.0	0	0.0	100.0
80	0	0.0	100.0	0	0.0	100.0
	Mean		36.05	Mean		38.09
	SD		11.20	SD		11.28
	Coef. Ske	wness	.57		Skewness	.67
		tosis	26		Kurtosis	27

Table 3
Univariate Distributions for AIM Scales

		Englis	<u>h</u>		Music	
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pct.
0 1 2 3 4 5 6 7 8 9 10 11 2 13 4 15 6 17 8 19 20 1 22 23 4 5 6 27 8 29 30 31 32	289 273469 4955756466666666666666666666666666666666	1122222333333333443333333223222212	1.6 0 2 2 0 7 7 7 1 6 4 1 0 9 8 9 0 8 7 3 0 4 8 8 7 4 4 0 5 6 8 4 0 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	510 5391 414 5506 5778 6858 648 648 648 649 654 654 654 654 654 654 654 654	002450744508288594603976623028282 323233333343433333333222222221213	3.5.8.26.1.1.9.1.0.7.2.1.5.0.1.4.3.0.6.3.5.8.0.8.0.8.0.1.1.8.26.1.1.9.1.0.7.2.1.5.0.1.4.3.0.6.3.5.8.7.0.8.0.8.0.8.0.8.0.8.0.8.0.8.0.8.0.8
	Mean SD Coef. Coef.	Skewnes Kurtosi			. Skewness	

Table 3 (Continued)

	Soc	ial Scie	ence		Math	<u>1</u>
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pct.
0 1 2 3 4 5 6 7 8 9 0 11 12 13 14 15 16 17 18 19 20 21 22 22 24 25 26 27 28 29 30 31 31 31 31 31 31 31 31 31 31 31 31 31	347 343 343 454 450 450 450 450 450 450 450 450 450	0.70.36.76.68.9.1.1.30.8.1.08.9.3.4.2.1.8.8.5.9.8.0.7.5.3.3 2.2.2.2.2.2.3.3.3.3.4.4.3.3.3.3.2.2.2.2.	2.7.7.0.5.2.8.5.3.1.3.4.8.8.6.7.7.5.4.7.2.3.5.3.1.6.4.3.2.0.5.7.0.3.3.3.4.4.8.5.5.5.5.3.1.6.4.3.2.0.5.7.0.3.3.3.4.4.5.5.5.5.5.6.6.7.7.7.8.8.8.8.9.5.0.5.7.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0	673 470 756 756 750 561 563 569 563 541 508 547 508 547 508 543 543 543 543 544 549 694 694	0.843459491372928008836344222222222 424343333332323232322222222	4.6.7.2.4.9.4.3.7.6.6.0.7.9.8.0.8.8.8.6.4.7.3.6.0.4.6.9.2.8.2.3.9.0.4.5.8.8.8.6.4.7.3.6.0.4.6.9.2.8.2.3.9.0.0.0.4.6.9.2.8.2.3.9.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0.0
	Mean SD Coef. Coef.	Skewne Kurtos	16.95 9.06 ess04 eis -1.04	•	Skewness Kurtosis	14.80 9.76 .22 -1.18

Table 3 (Continued)

	Home Ec	onomics		<u>Fi</u>	ne Arts	
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pct.
0 1 2 3 4 5 6 7 8 9 10 11 2 13 4 15 6 17 8 9 20 21 2 23 4 25 6 7 8 29 30 12 32 25 26 7 8 29 30 31 2	276 228 332 369 391 374 371 458 561 459 5515 495 561 597 519 561 581 591 591 591 591 591 591 591 591 591 59	1.30728300232670365501799326578817	1.6 3.4 6.7 10.3 14.8 10.3 14.8 10.3 10.3 10.4 10.3 10.4 10.3 10.4 10.3 10.4 10.3 10.4 10.3 10.	173 197 197 197 197 199 199 199 199 199 199	1.5811579874766707161884523187952	1.0 2.7 5.6 7.2 9.7 9.2 14.7.5 14.7 17.5 14.7 17.5 14.7 17.5 14.7 17.8 16.4 18.3 19.8 19.9 19.0 19.0 19.0 19.0 19.0 19.0 19.0
		Skewness Kurtosis			Skewness Kurtosis	

Table 3 (Continued)

	Physic	al Scien	ce		Engineeri	ng
Score	Freq.	Pct.	C-Pet.	Fre	q. Pct.	C-Pet.
0 1 2 3 4 5 6 7 8 9 10 112 134 15 6 7 8 9 21 22 324 526 7 8 9 30 132 14 5 6 7 8 29 30 31 32	8994 5965 5965 5965 5965 5965 5965 5965 5	335716170899071092190765847365092 3323232323232322222222325	5.5 8.9 14.9 17.0 18.9 19.5 19.5 19.5 19.5 19.5 19.5 19.5 19	47 49 50 57 49 62 57	2.5 3.8 3.8 3.8 3.8 3.8 3.8 3.8 3.8 3.8 3.8	36.9.1.4.1.1.0.9.5.3.0.7.4.3.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.9.8.6.5.4.8.7.4.8.0.2.2.6.5.2.2.2.2.2.2.2.2.2.2.2.2.2.2.2.2
	Mean SD Coef. Coef.	Skewnes Kurtosi		S: C	ean D oef. Skewnes oef. Kurtos:	

Table 3 (Continued)

	Biolog:	ical Scier	nces	Secretarial					
Score	Freq.	Pet.	C-Pct.	Freq.	Pct.	C-Pct.			
0 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 24 25 26 27 28 29 30 31 32	234354 45156 45156 45156 4515 4515 4515 4515	1.40167982247860140858440190684408	1.8 8 9.5 2 1 9.0 2 7 3 1 7 7 9 3 2 0 5 3 6 0 0 1 9 0 6 4 8 2 2 2 3 1 5 1 7 6 1 3 2 0 5 3 6 6 6 7 7 6 1 9 9 9 9 9 7 1 0 0 6 4 8 2 2 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	808 2019 2194 2338 3348 3348 431 431 431 431 431 431 431 431 431 431	0.6.2.3.4.6.0.0.4.5.8.6.9.9.4.6.8.8.8.7.0.4.7.2.4.3.5.6.8.1.3.8.3.3.3.3.3.3.3.3.3.3.3.3.3.3.3.3	0.5 1.3 3.6 0.6 10.7 13.6 10.7 13.6 10.7 13.6 10.7 13.6 10.7 13.6 13.7 13.7 13.7 13.7 13.7 14.7 15.8 14.9 15.9 16.6 16.6 17.7 17.6 16.6 16.6 16.6 16.6			
		Skewness Kurtosis	16.40 8.53 1.90 96		Skewness Kurtosis				

Table 3 (Continued)

Control of the second of the s

	Fore	eign Lan	guages		Executi	ve
Score	Freq.	Pct.	C-Pct.	Freq.	Pct.	C-Pct.
0 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26 27 28 29 30 31 32	982 439 504 369 315 311 368 369 491 492 494 513 458 458 458 513 513 612 578 612 578 612 613 614 615 615 616 617 617 618 618 619 619 619 619 619 619 619 619	860228981922519961076707073066545 2211121222223222232323333439	5.8 8.4 11.3 13.5 15.7 17.6 19.4 21.3 23.3 27.5 29.6 32.1 29.5 29	102 68 168 139 230 232 336 351 486 512 575 679 630 743 777 7812 683 719 681 571 681 571 681 571 683 571 683 571 683 571 683 571 683 571 683 571 683 571 683 571 683 571 683 571 683 571 683 683 683 683 683 683 683 683 683 683	0.0000	0.6 0.0 1.2 2.4 5.7 9.1 1.5 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0
		Skewness Kurtosis			Skewness Kurtosis	

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PSAT-V and PSAT-Q Bivariate Distribution

TOTAL		0	7	45	32	142	146	331	390	61 4	161	648	1058	1074	1202	1554	1484	1947	2064	2016	979	276	17001
84						Н																	-
76 77				ო		Н	Н	7			Н												۵
73						Q	Q			Н	Н	н											7
72			Н	*	Q	4	Ø	۲	ო	ო	ო		a	Н									S S
69			н	ω		13	80	7	70	σ	5	ς.	ო	႕		Н							75
3 9				9	ณ	11	14	16	נו	17	ω	σ	ณ	a		a							706
61 63				9	ដ	17	15	27	25	34	30	13	378	ω	<u>~</u>	a	႕						219
929			a	۲-	‡	ส	17	9	30	נל	84	1 47	25	17	₹	9	Н						305
55 57			٦	#	€	Ŕ	16	33	04	₹	26	745	04	33	ส	13	5	'n	ო				405
52 54			αı	7	ന	†T	83	52	88	76	76	16	93	47	77	23	8	9	Ŋ	α			4 49
PSAT-V 46 49 48 51				#	a	5	16	9	89	81	100	86	66	87	72	70	33	11	Ħ	#	႕		802
8,58					a	13	ដ	31	37	82	97	102	127	95	96	77	63	30	12	ω	႕		886
43 45				Н	a	7	√	31	37	8	נננ	120	158	156	138	133	105	62	52	16	ന		1206
3 <u>3</u>				႕		Н	9	51	72	55	26	103	145	165	176	198	328	108	78	88	ω	4	1339
37						႕	Н	य	ቪ	33	77	72	127	131	186	212	192	189	106	77	17	7	1455
38						Н	a	ω	12	%	33	8	96	136	166	228	211	237	220	136	‡	_	1637 1455
33 33								4	4	ω	32	52	[]	103	142	242	251	329	306	242	81	12	1879
କ୍ଷ ନ								ന	a	임	נו	13	200	44	78	135	191	257	549	251	86	56	1368
23									Н	4	m	12	16	37	51	120	165	303	402	414	184	62	1774
22 24 24											a		ſV	ω	13	62	88	216	277	396	210	23	1334
8 5												႕	Н	m	0/	58	9	194	345	844	332	101	1520
	PSAT-Q	79-80	76-78	73-75	70-72	69-19	99-49	61-63	58–60	55-57	52-54	49-51	94-94	43-45	40-42	37-39	34-36	31-33	28-30	25-27	22-24	20-21	TOTAL

Table 5

Approximate Means and Standard Deviations for PSAT-V and PSAT-Q

PSA'T-V	PSA*	T-Q	PSAT-Q	PSA	T–V
Interval	Mean	SD	Interval	Mean	SD
20-21	27.1	4.3	20-21	24.5	4.9
22-24	28.4	4.9	22-24	25.0	5.2
25 - 27	30.1	6.0	25-27	26.7	5.9
28-30	32.3	7.0	28-30	28.8	6.9
31-33	34.4	7.4	31-33	30.8	7.1
34-36	36.9	8.2	34-36	34.3	7.7
37-39	39.4	8.5	37-39	36.5	7.9
40-45	42.0	8.5	10-15	39.1	7.7
43-45	44.4	8.8	43-45	41.1	8.2
46-48	46.9	8.8	46-48	43.4	8.2
49-51	48.7	8.7	49-51	45.3	8.7
52-54	51.4	8.6	52 - 54	46.8	8.6
55-57	53.2	9.1	55 - 57	49.1	8.8
58-60	55 . 6	8.0	58-60	50.9	8.3
61-63	57.0	8.3	61-63	52.3	9.0
64-66	60.2	7.7	64-66	55.7	8.1
67-69	61.2	8.2	67-69	58.0	8.1
70-72	62.3	8.4	70 - 72	57.8	6.7
73 - 75	60.7	7.0	73-75	62.1	8.1
76-78	67.1	7.3	76-78	59.9	6.6
79-80	68.0		79 – 80	400 Miles dates dates	

Table 6

PSAT-V and PSAT-Q Bivariate Distribution for Grouped Data

	C-PCT		5.4	10.0	19.6	29.9	39.3	51.0	6.09	80.7	90.3	100.0		
	TOTAL		426	783	0491	1743	1591	1989	1692	3368	1627	1644	17001	
	57-80		384	169	203	100	35	19	4	ч			915	5.4
	52-56		203	143	233	747	85	04	19	1,4	ત્ય		988	10.6
	46-51		185	209	397	349	227	179	42	61	Ħ	m	1688	20.5
PSAT-V	42-45		4 ^L	120			280	442	138	122	15	13	1682	30.4
PS	38-41		39	65	197	254	254	280	199	212	T †	25	1566	39.6
	35-37		덚	94	139	238	276	335	264	347	111	59	1836	50.4
	29-34		1.8	22	130	218	313	519	984	898	354	273	3231	4.69
	26–28			ω	17	9	85	206	260	469	332	356	2018	81.3
	22-25			Н	9	57	56	119	158	539	389	904	1659	91.1
	20-21				Н	Μ	10	£4	85	764	372	509	1520	100,0
		PSAT-Q	59–80	55-58	49-54	84-44	40-43	36-39	33-35	28-32	26-27	20-25	TOTAL	C-PCt

Table 7
PSAT-V and PSAT-Q Means and Standard Deviations for Grouped Data

PSAT-V	PSA'	T–Q	PSAT-Q	PS/	VT/
Interval	Mean	SD	Interval	Mean	SD
20-21	27.2	4.1	2025	25.6	5.1
22-25	28.7	5.0	26-27	26.9	5.7
26-28	30.6	6.1	28-32	29.4	6.8
29-34	34.2	7.5	33 - 35	33.3	7.5
35-37	38.0	8.2	36-39	36.0	8.3
38-41	40.7	8.6	40-43	39.7	8.6
42-45	44.1	8.6	44-48	42.9	9.6
46-51	47.8	8.8	49-54	47.0	10.5
52-56	51.8	8.8	55- 58	50.8	10.9
57-80	56.9	8.6	59 – 80	56.6	11.0

Table 8

Names and Number of Items for Each Scale in Army Classification Battery

Scale Name	Number of Items
Verbal Test	50
Arithmetic Reasoning	40
Pattern Analysis	50
Mechanical Aptitude	45
Army Clerical Speed	110
Army Radio Code	150
Shop Mechanics	40
Automotive Information	40
Electronics Information	40
Classification Inventory	125
General Information Test	50

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Table 9

Names and Number of Items for Each Scale in Army Differential MOS Battery

BARTON CONTRACTOR CONTRACTOR OF LIMITED STATES OF A LOWER OF A STATES TO SERVICE OF THE SERVICE

Scale Name	Number of	Items
Subtraction and Division Test	100	
Tool Knowledge Test	20	
Electronics Interest	20	
Mechanical Interest	20	
Clerical Interest	20	
General Adjustment	20	
Electronics Knowledge Test	20	
Mechanical Principles	20	
Mathematical Knowledge Test	20	
Science Knowledge Test	20	
Pattern Analysis Test	20	
Bio-Chem Information Test	30	
Electronics Pictures Test	20	

Table 10

Means, Standard Deviations and Actual Score Range for ACB

and Differential MOS Scales

<u>Mean</u>	<u>SD</u>	Actual Score Range
108.9	18.6	50 - 152
106.1	18.8	50 - 160
105.6	20.3	57 - 155
107.1	17.0	40-160
109.7	17.2	50-150
99.3	27.0	50 – 150
106.5	17.2	39 - 154
106.9	18.7	55 - 150
106.4	18.8	40-160
101.2	19.1	40-160
103.0	16.9	59-160
37.0	14.8	0-100
12.5	4.0	0-20
7.7	3.9	0-20
11.7	3.8	0-20
11.2	3.1	0-20
13.0	3.1	0-20
9.8	3.7	0-20
9.6	3.6	0-20
9.1	4.1	0-20
13.0	5.4	0-30
11.6	4.0	0-20
17.6	5.9	0-30
9.9	4.2	0-20
	108.9 106.1 105.6 107.1 109.7 99.3 106.5 106.9 106.4 101.2 103.0 37.0 12.5 7.7 11.7 11.2 13.0 9.8 9.6 9.1 13.0 11.6 17.6	108.9 18.6 106.1 18.8 105.6 20.3 107.1 17.0 109.7 17.2 99.3 27.0 106.5 17.2 106.9 18.7 106.4 18.8 101.2 19.1 103.0 16.9 37.0 14.8 12.5 4.0 7.7 3.9 11.7 3.8 11.2 3.1 13.0 3.1 9.8 3.7 9.6 3.6 9.1 4.1 13.0 5.4 11.6 4.0 17.6 5.9

Table 11

ACB-V and ACB-A Bivariate Distribution for Grouped Data

	C-PCT	5.0	10.7	16.4	22.1	28.6	34.3	38.9	44.3	50.5	55.6	63.2	72.7	32.5	9.06	100.0		
	TOTAL	1115	127,1	1255	1271	1437	1266	1021	9611	1367	1119	1682	2098	2174	1784	2076	22132	
	137	299	166	118	66	58	38	23	33	27	11	ηг	17	בן	м	ณ	929	2.4
	133	295	210	1.74	147	149	8	51	13	††	31	31	31	22	ľ	t-	1328	10.2
	130	741	237	991	141	114	97	55	65	75	42	53	39	33	15	15	1297	16.1
	126 129	82	125	191	95	131	104	19	ħΤ	78	99	52	45	33	7₹	22	1132	21.2
	123	8	152	122	756	115	98	37	85	103	31	85	96	7,2	55	50	1242	26.8
	119	54	20	76	111	126	104	75	110	2	61	8	100	84	36	7,7	1162	32.0
<u>۲</u>	117	32	85	97	110	110	96	20	72	8	25	93	83	65	7 ⁷	31	1133	37.1
ACB	114 116 11	36	70	76	98	151	163	911	110	134	8	118	125	107	η9	89	1546	1.44
	111	18	23	56	75	75	917	27	83	65	36	63	81	28	50	28	751	47.5
•	108	15	22	712	64	78	87	98	88	100	82	109	140	123	42	85	1201	52.9
	103	27	52	72	81	ננו	129	127	149	164	136	245	549	242	176	161	2151	62.6
•	102	10	59	5.5	35	16	83	77	88	911	117	193	226	260	223	216	1797	70.7
;	91	7	17	53	148	L 9	88	66	110	130	148	506	323	386	324	601	2391	81.5
5	90	Q	11	21	76	36	43	99	55	113	121	196	293	375	373	66il	2220	91.5
(9, 89 93, 93	Н	N	9	5	22	22	33	64	89	81	125	250	369	348	بوک	1852	100.0
	ACB-A	135-160	130-134	126-129	121-125	118-120	115-117	113-114	109-112	107-108	105-106	100-104	66-55	89 - 94	81-88	39-80	TOTAL	C-PCT

Table 12

ACB-V and ACB-A Means and Standard Deviations for Grouped Data

ACB-V	ACE	-A	ACB-A	ACI	B V
Interval	Mean	SD	Interval	Mean	SD
39 - 83	89.2	14.8	39-80	92.4	14.8
84-90	92.3	14.5	81-88	95.3	14.4
91-97	95.6	15.5	89-94	97.5	15.7
98-102	98.6	15.6	95 - 99	101.7	16.0
103-107	102.7	15.8	100-104	105.2	15.5
108-110	104.4	14.9	105-106	106.0	15.6
111-113	107.6	15.4	107-108	109.4	15.8
114-116	109.5	14.8	109-112	111.5	14.9
117-118	111.9	14.7	113-114	111.0	14.9
119-122	113.2	13.7	115-117	114.8	14.2
123-125	116.2	15.0	118-120	117.3	14.3
126-129	117.0	14.5	121-125	121.3	13.2
130-132	120.6	14.3	126-129	123.2	13.3
133-136	124.7	13.9	130-134	126.3	12.4
137-160	128.5	14.4	135-160	131.8	12.1

Table 13
Selected Comparisons of Extrapolated Correlations with
Applicant Group Values for the PSAT-AIM Data
with Selection Based on PSAT-V

Variable Correlated	Correlation	Corre	lation	s in S	electe	Extrapolated Correlations					
with Selection	in Applicant	Per	cent i	n Sele	cted G	Percent in Selected Group					
Variable	Group	10	30	50	70	90	10	30	50	70	90
PSAT-Q	.75	.36	.49	.58	.65	.73	.66	.68	.71	.73	.75
English	.27	.11	.19	. 24	.27	.28	.24	.30	.32	.33	. 30
Music	.11	.02	.08	.11	.13	.13	.04	.14	.16	.16	.13
Soc. Sci.	.22	.11	.17	.20	.23	. 23	.25	.28	.28	.28	.25
Math.	.15	.ივ	.11	. 14	.15	.16	.17	.17	.19	.19	.17
Phy. Sci.	. 14	.07	.14	.16	.17	.16	.16	.23	.22	.21	.17
Engr.	02	00	.00	.00	.00	01	01	01	.01	.00	.01
H. Econ.	12	13	12	14	13	11	27	20	19	16	12
Fine Arts	.06	04	.02	.02	• 04	.06	09	.03	.03	.05	.06
Bio. Sci.	.05	.01	.04	.07	.07	.07	.01	.07	.09	.08	.07
Secretar.	-·2 ¹ +	15	23	24	27	25	31	36	33	32	26
For. Lang.	.22	.10	.15	.18	.21	.22	.22	. 24	. 24	.26	. 24
Exect.	08	09	09	10	10	09	20	15	14	12	09 ·

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Table 14
Selected Comparisons of Extrapolated Correlations with
Applicant Group Values for the PSAT-AIM Data
with Selection Based on PSAT-V

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.00	.06		ļ	30	50	70	00
01		.11	0.7			1 3	90
	00		101	19	.12	.14	.13
	.02	.03	06	10	.05	.06	.05
.07	.10	.13	.10	08	.15	.16	.15
.47	.45	.41	.48	.49	.47	.45	.41
. 24	. 24	.22	.25	.18	.28	.27	.23
.17	.15	.13	.15	.19	.15	. 14	.12
21	21	19	31	16	24	23	20
08	06	03	20	15	06	04	02
.07	.07	.07	.04	• 04	.09	.09	.07
21	24	24	50	06	29	29	25
.03	.07	.11	.05	10	.11	.13	.12
- 05	- .06	06		.04			
	.17 21 08 .07 21	.17 .15 2121 0806 .07 .07 2124 .03 .07	.17 .15 .13212119080603 .07 .07 .07212424 .03 .07 .11	.17 .15 .13 .15 21211931 08060320 .07 .07 .07 .04 21242420 .03 .07 .11 .05	.17 .15 .13 .15 .1921211931160806032015 .07 .07 .07 .04 .042124242006 .03 .07 .11 .0510	.17 .15 .13 .15 .19 .15212119311624080603201506 .07 .07 .07 .04 .04 .09212424200629 .03 .07 .11 .0510 .11	.17 .15 .13 .15 .19 .15 .14 21211931162423 08060320150604 .07 .07 .07 .04 .04 .09 .09 21242420062929 .03 .07 .11 .0510 .11 .13

Table 15

Selected Comparisons of Extrapolated Correlations with Applicant

Group Values for the Army Data with Selection Based on ACB-V

Variable Correlated	Correlation	Correlations i			Sele	cted	Group	Extrapolated Correlations				ns	
with Selection	in Applicant	Pe	rcent	in Se	lecte	d Gro	up	Pe	rcent	in Se	lecte	d Gro	up
Variable	Group	10	20	40	60	80	90	10	20	40	60	80	90
ACB-A	.60	.16	.27	.38	.47	•55	.58	.42	.60	.64	.65	.65	.64
Pat. Anal.	.40	.09	.14	.23	.30	•37	.40	.26	.36	.43	.46	.46	.45
Mech. Apt.	.45	.10	.16	. 26	.33	.40	.43	.27	•39	.48	.50	.49	.48
ACS	.36	.07	.10	.17	• 24	.30	.33	.19	.25	.32	. 37	.38	.38
Shop M	.31	.01	.06	.11	.18	.25	.28	.04	.15	.22	.29	.32	.33
Auto Inf.	.23	03	01	.03	.10	.16	.20	09	01	.06	.16	.21	.23
EL Inf.	.40	.08	.12	.20	.28	.35	•37	.22	.30	.39	.43	.43	.42

Table 16
Selected Comparisons of Extrapolated Correlations with Applicant
Group Values for the Army Data with Selection Based on ACB-V

Correlations	Correlation	Corr	elati	ons i	n Sel	ected	Group	Ex	trapo	lated	Corr	elati	ons
between ACB-A and	in Applicant	Pe	rcent	in S	elect	ed Gr	oup	Pe	rcent	in S	elect	ed Gr	oup
	Group	10	20	40	60	80	90	1.0	20	40	60	80	90
Pat. Anal.	.51	.42	.45	.46	.49	.51	.51	.47	.54	•57	•57	.56	.54
Mech. Apt.	.48	.32	.37	.39	.42	.45	.47	.38	.49	•53	•53	. 52	.51
ACS	.46	.38	•39	.38	.41	•##	.45	.41	. 1414	.46	.48	.49	.48
Shop M	•32	.17	.21	.22	.26	.29	.30	.17	.25	.29	•33	.34	.34
Auto Inf.	.26	.10	.14	.16	.19	.22	. 24	.06	.10	.16	.23	.26	.26
EL Inf.	.40	. 24	.28	.30	.33	.36	.38	.30	.38	.42	.1414	.43	.42

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Table 17
Comparison of Two New Correction Methods with
Standard Method Using PSAT Data

	,,				
	Percent in Selected Group				
	10	30	50	70	90
PSAT-Q	·				
Applicant Group Correlation Standard Method New Method 1 New Method 2	.75 .66 .78 .64	.75 .68 .78 .72	.75 .71 .77 .75		.75 .75 .76
English					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.27 .24 .23 .21	.27 .20 .29 .28		.33	.30 .30
Music					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.11 .04 .04 .04	.11 .1/: .14 .14	.16	.16	
Soc. Sci.					
Applicant Group Correlation Standard Method New Method 1 New Method 2		.22 .28 .28 .27	. 28 . 28	. 28 . 28	.25
Math	ļ				
Applicant Group Correlation Standard Method New Method 1 New Method 2	.15 .17 .17 .16	.15 .17 .17 .17		.19	.17
Phy. Sci.	•				
Applicant Group Correlation Standard Method New Method 1 New Method 2	.14 .16 .15 .15	.14 .23 .22 .22	.14 .22 .22 .22	.21	.11. .17 .17

Table 17 (continued)

•	Perc	ent in	Selec	ted Gr	Oil D
	10	30	50	70	90
Engr.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	02 01 01	0l 0l	.01		01 01
H. Econ.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	12 27 28 27	20	19	16 16	12 12
Fine Arts					
Applicant Group Correlation Standard Method New Method 1 New Method 2		.06 .03 .03 .03	.06 .03 .03 .03		.06
Bio. Sci.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.05 .01 .01	.07	.05 .09 .09	.05 .08 .08 .08	.05 .07 .07
Secretar.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	31	36 36	33 33	24 32 32 32	26 26
For. Lang.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.22 .22 .20 .18	.22 .24 .23 .22	. 22 . 24 . 24 . 23	.25	
Exect.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	20	15 15	14 14	08 12 12 13	09 09

Table 18

Comparison of Two New Correction Methods with

Standard Method Using Army Data

	Ι "				
	Pero 20	ent ir 40	60 60	eted G	roup 90
	20				
ACB-A		_	_	_	4
Applicant Group Correlation Standard Method New Method 1 New Method 2	.60 .60 .58	.60 .64 .60 .73	.65	.65 .61	
Pat. Anal.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.40 .36 .34 .37	.40 .43 .41 .45	.46	.46 .44	
Mech. Apt.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.45 .39 .46 .44	.45 .48 .49		.49 .48	.48
ACS					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.36 .25 .28 .26	.36 .32 .31 .33	. 37	.38	.38
Shop M					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.31 .15 .16	.22	.31 .29 .29 .31	.32	.33
Auto. Inf.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.23 01 02 02	.23 .06 .06 .06	.16	.23 .21 .21 .22	. 25 . 23 . 23 . 23
El. Inf.					
Applicant Group Correlation Standard Method New Method 1 New Method 2	.40 .30 .41 .35	.40 .39 .42 .44	.40 .43 .41 .48	.40 .43 .42 .47	.40 .42 .42 .45

Table A Univariate Distribution for ACB-V Scale

ACB-Verbal

Score	Freq.	Pct.	C-Pet
50	72	0.3	0.3
51	0	0.0	0.3
52	Ō	0.0	0.3
53	ŏ	0.0	
54	0		0.3
55		0.0	0.3
56	25	0.1	0.4
	3 2	0.0	0.5
57		0.0	0.5
58	0	0.0	0.5
59	0	0.0	0.5
60	31	0.1	0.6
61	0	0.0	0.6
62	19	0.1	0.7
63	0	0.0	0.7
64	20	0.1	0.8
65	28	0.1	0.9
66	1	0.0	0.9
67	31	0.1	1.0
68	77	0.3	1.4
69	30	0.1	1.5
70	37	0.2	1.7
71	119	0.5	2.2
72	53	0.2	2.5
73	50	0.2	2.7
74	128	0.6	3.3
75	50	0.2	
76	59	0.3	3.5
77	119		3.8
		0.5	4.3
78 70	65 7 3	0.3	4.6
79	73	0.3	4.9
80	362	1.6	6.6
81	3	0.0	6.6
82	394	1.8	8.3
83	, 1	0.0	8.4
34	409	1.8	10.2
85	110	0.5	10.7
86	259	1.2	11.9
87	134	0.6	12.5
88	609	2.7	15.2
89	5	0.0	15.2
90	694	3.1	18.4
91	155	0.7	19.1
92	657	3.0	22.0
93	5	0.0	22.1
94	507	2.3	24.3
95	137	0.6	25.0
96	±31 74	0.3	
90 97	859		25.3
91 98		3.9	29.2
	257 200	1.2	30.3
99	200	0.9	31.2
100	1149	5.2	36.4

Table A (Continued)

ACB-Verbal

Score	Freq.	Pct.	C-Pct.
101	2	0.0	36.4
102	190	0.9	37.3
103	988	4.5	41.7
104	5	0.0	41.8
105	256	1.2	42.9
106	679	3.1	46.0
107	227	1.0	47.0
108	5	0.1	47.0
109	245	1.1	48.1
110	952	4.3	52.4
111	308	1.4	53.8
112	142	0.6	54.4
113	302	1.4	55.8
114	1011	4.6	60.4
115	151	0.7	61.0
116	387	1.7	62.8
117	301	1.4	64.1
118	833	3.8	67.9
119	449	2.0	69.9
120	16	0.1	70.0
121	554	2.5	72.5
122	144	0.6	73.2
123	512	2.3	75.5
124	3	0.0	75.5
125	733	3.3	78.8
126	716	3.2	82.0
127	8	0.0	82.0
128	407	1.8	83.9
129	2	0.0	83.9
130	1298	5.9	89.7
131	3	0.0	89.8
132	3	0.0	89.8 93.0
133	710	3.2 0.0	93.0
134	1	0.0	93.0
135	621		
136		2.8 0.0	95.8 95.8
137	3 6	0.0	95.8
138 139	2	0.0	95.8
140	275	1.2	97.1
140	1	0.0	97.1
141	0	0.0	97.1
11,3	ŭ	0.0	97.1
144	0	0.0	97.1
145	364	1.6	98.7
/	_	-	- •

Table A (Continued)

ACB-Verbal

Score	Freq.	Pct.	C-Pct.
146	1	0.0	98.7
147	1	0.0	98.8
148	1	0.0	98.8
149	0	0.0	98.8
150	1	0.0	98.8
151	0	0.0	98.8
152	276	1.2	100.0

Mean		108.89
SD		18.61
Coef	Skewness	14
Coef	Kurtosis	36

Table B
Univariate Distribution for ACB-A Scale

ACB-Arithmetic

Score	Freq.	Pct.	C-Pct.
50 50 50 51 52 53 54 55 55 57 58 59 60 61 62 63 64 65 66 67 68 69 70 70 70 70 70 70 70 70 70 70 70 70 70	121 00 00 27 00 32 04 113 14 00 90 95 10 130 218 137 91 147 140 140 140 140 140 140 140 140 140 140	0.5 0.0 0.0 0.0 0.1 0.0 0.0 0.0 0.0 0.0 0.0	0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5

Table B (Continued)

ACB-Arithmetic

Score	Freq.	Pct.	C-Pct.
91	303	1.4	22.8
92	72	0.3	23.1
93	882	4.0	27.1
94	36	9.2	27.2 28.3
95 96	245 675	1.1 3.0	31.4
90 97	366	1.7	33.0
98	141	0.6	33.7
99	675	3.0	36.7
100	158	0.7	37.4
101	350 25.5	1.6	39.0 42.7
102	817 356	3.7 1.6	44.3
103 104	350 4	0.0	44.3
105	1110	5.0	49.3
106	10	0.0	49.4
107	573	2.6	51.9
108	797	3.6	55 • 5
109	9	0.0 1.7	55 . 6 57 . 3
110	385 463	2.1	59.4
111 112	34 1	1.5	60.9
113	9	0.0	61.0
114	1012	4.6	65.5
115	174	0.8	66.3
116	430	1.9	68.3 71.3
117	664 459	3.0 2.1	73.3
118 119	429 5	0.0	73.4
120	976	4.4	77.8
121	7	0.0	77.8
122	415	1.9	79.7
123	150	0.7	80.3
124	9	0.0	80.4 83.5
125	697 487	3.1 2.2	85.7
126 127	384	1.7	87.4
128	10	0.0	87.5
129	374	1.7	89.2
130	575	2.6	91,8
131	393	1.8	93.5

Table B (Continued)

ACB-Arithmetic

Score	Freq.	Pct	C-Pct.
132	0	0.0	93.5
133	3	0.0	93.6
134	312	1.4	95.0
135	3	0.0	95.0
136	294	1.1	96.1
137	2	0.0	96.1
138	188	0.8	97.0
139	1	0.0	97.0
140	3	0.0	97.0
141	207	0.9	97.9
142	0	0.0	97.9
143	153	0.7	98.6
144	Ţ	0.0	98.6
145	4	0.0	98.6
146	119	0.5	99.2
147	1	0.0	99.2
148	1	0.0	99.2
149	70	0.3	99.5
150	0	0.0	99.5
151	0	0.0	99.5
152	55	0.2	99.7
153	0	0.0	99.7
154	0	0.0	99.7
15 5	0	0.0	99.7
156	O	0.0	99.7
157	0	0.0	99.7
158	0	0.0	99.7
159	0	0.0	99.7
160	59	0.3	100.0

Mean		106.14
SD		18.78
Coef.	Skewness	15
Coef	Kurtosis	21

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The Pearson formulas for correcting correlation coefficients for restriction of range are based on crucial assumptions of linearity of regression and homoscedasticity of the error distributions. Some small previous studies have suggested that these formulas are reasonably accurate providing extreme selection is not involved. These studies tend to suggest that the formulas typically provide undercorrections in most instances. The present study involved two very large data sets and attempted to verify the accuracy or inaccuracy of these formulas and the assumptions on which they are based for both moderate and extreme selection. Generally, it was found that the linearity assumption was reasonably well satisfied except in the extreme tails of the distribution while the homoscedasticity assumption was not. In neither set of data distribution correction formulas work as well as previous research had led the authors to expect they would. Undoubtedly this was due to the invalidity of the homoscedasticity assumption. Some methods for taking into account heteroscedasticity of errors were studied and some very minor improvements were found. However, no method seems to have any general validity.

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